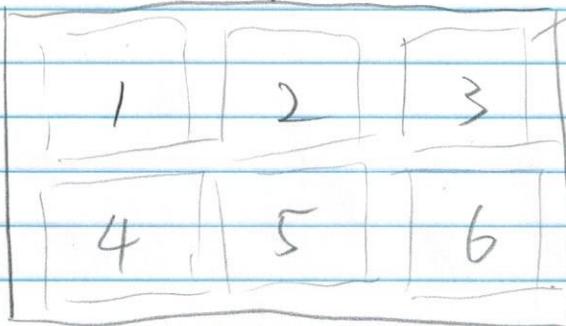


①

L3 Supp. Material.

Fair Dice example.



Sample space. Ω

$$\{1, 2, 3, 4, 5, 6\}$$

Event space. $\left\{ \{1\}, \{2\}, \dots, \{1, 2\}, \{1, 3\}, \dots, \{1, 2, \dots, 5, 6\}, \{\emptyset\} \right\}$

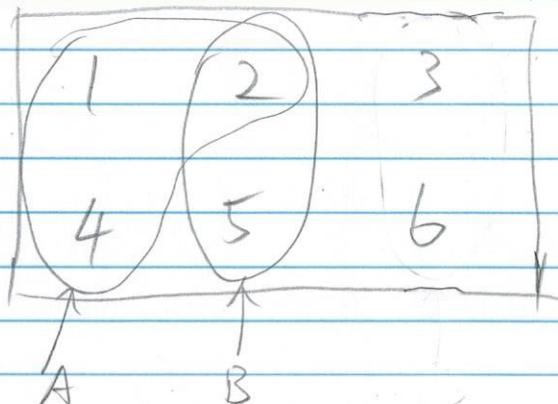
F

Probability measure:

$$\begin{aligned} \{1\} &\xrightarrow{P} \frac{1}{6} \\ \{2\} &\xrightarrow{P} \frac{1}{6} \\ \{6\} &\xrightarrow{P} \frac{1}{6} \end{aligned} \quad \left. \right\} P$$

(2)

Fair Dice Cont'd



$$P(A) = \frac{3}{6} = \frac{1}{2}$$

$$P(B) = \frac{2}{6} = \frac{1}{3}$$

$$P(A \cup B) = \frac{4}{6} = \frac{2}{3}$$

$$< P(A) + P(B)$$

$$A \cup B = \{1, 2, 4, 5\}$$

$$P(A \cap B) = \frac{4}{6} = \frac{2}{3}$$

$$= 1 - P(B)$$

$$A \cap B = \{2\}$$

$$A \cap B = \{1, 3, 4, 6\}$$

$$P(A \cap B) = \frac{1}{6}$$

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{\frac{1}{6}}{\frac{1}{3}} = \frac{1}{2}$$



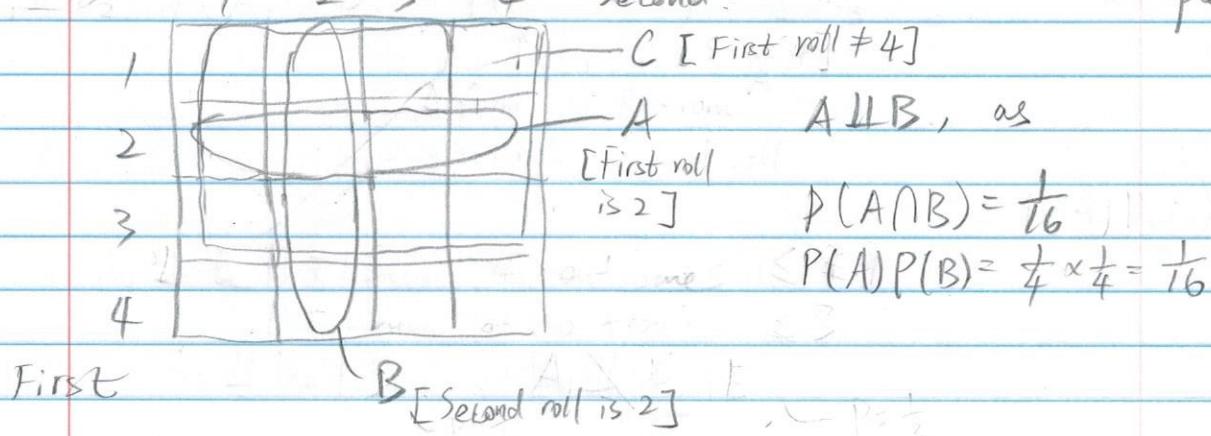
Independence

4-face dice, twice:

B as new

Sample space

A, B, C, D | 1 2 3 4 Second



(3)

1. $A \cap B | C$ as

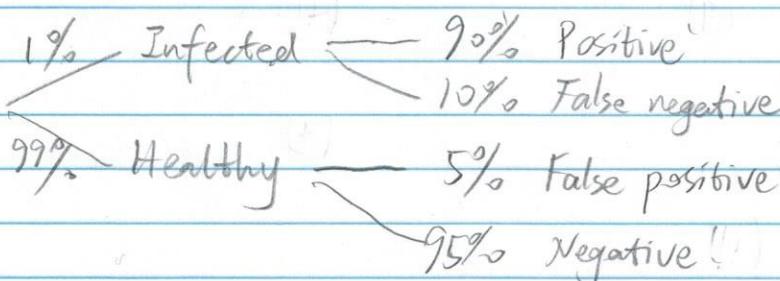
$$P(A \cap B | C) = \frac{P(A \cap B \cap C)}{P(C)} = \frac{\frac{1}{16}}{\frac{12}{16}} = \frac{1}{12}$$

$$P(A|C) P(B|C) = \frac{4}{12} \cdot \frac{3}{12} = \frac{1}{12}.$$

Let $D = [\text{Sum of two rolls} \leq 4]$ $A \cap B | D$

Bayes' Rule.

Screen test.

 B_1 - Infected, B_2 - Healthy A_1 - Positive, A_2 - Negative

$$P(B_1 | A_1) = \frac{P(A_1 | B_1) P(B_1)}{P(A_1)}$$

$$= \frac{0.9 \cdot 0.01}{0.058}$$

$$\approx 15.4\%$$

$$P(A_1) = P(A_1 | B_1)P(B_1) + P(A_1 | B_2)P(B_2)$$

$$= 0.9 \cdot 0.01 + 0.05 \cdot 0.99$$

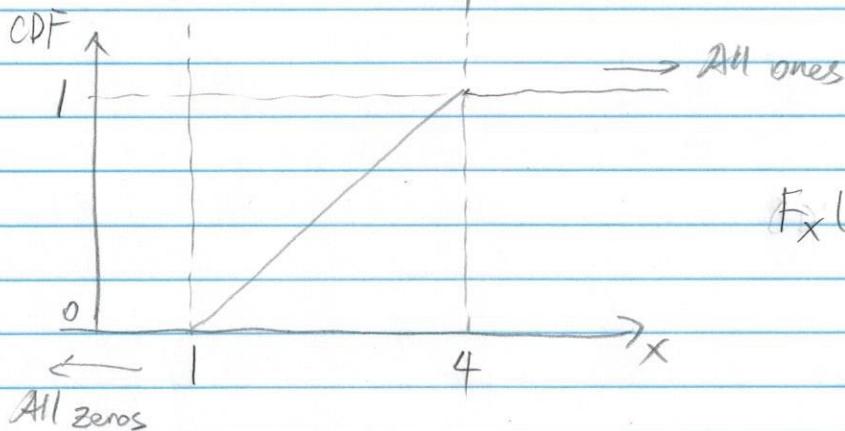
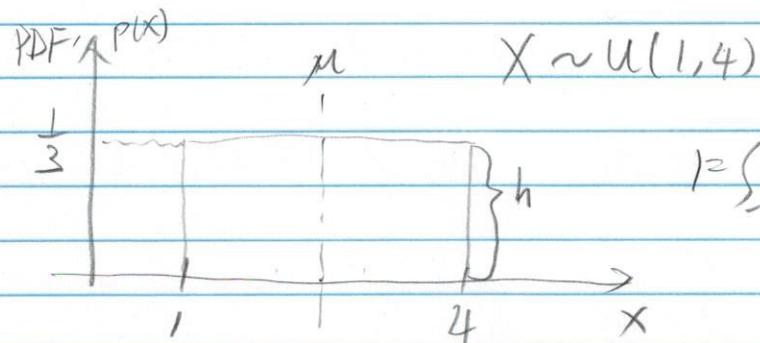
$$\approx 0.058$$

Positive test means 84.6% infection.

(4)

Example.

Uniform dist.



$$P(2 < x \leq 3) = \int_2^3 p(x) dx = F_x(3) - F_x(2) = \frac{1}{3}$$

Expectation:

$$E[X] = \int_{-\infty}^{\infty} x p(x) dx = \int_1^4 x \cdot \frac{1}{3} dx = \frac{5}{2}$$

Variance:

$$E[(X - E[X])^2] = \int_1^4 \frac{1}{3} (x - \frac{5}{2})^2 dx = \frac{3}{4}$$

Alt.

$$E[X^2] = \int_1^4 \frac{1}{3} x^2 dx = \frac{7}{2}$$

$$\text{Var} = E[X^2] - E[X]^2 = \frac{7}{2} - (\frac{5}{2})^2 = \frac{3}{4}$$

(5)

Covariance

$$\begin{aligned}\text{Cov}[X, Y] &= E[(X - EX)(Y - EY)] \\ &= E[XY - (EX)Y - X(EY) + (EX)(EY)] \\ &= E[XY] - (EX)(EY) - \cancel{(EX)(EY)} + \cancel{(EX)(EY)}\end{aligned}$$

$$\begin{aligned}\text{Var}[X+Y] &= E[(X+Y - EX - EY)^2] \\ &= E[(X-EX)^2 + (Y-EY)^2 + 2(X-EX)(Y-EY)] \\ &= \text{Var}[X] + \text{Var}[Y] + 2\text{Cov}[X, Y]\end{aligned}$$

Likelihood example

Since i.i.d.

$$P(D|w) = P(\{x_1, x_2, \dots, x_N\}|w) \quad \text{Same } w, \text{ since i.d.}$$

$$= P(x_1|w) P(x_2|w) \cdots P(x_N|w) \quad \text{Decompose, since}$$

$$= \prod P(x_i|w), \quad w = \{\mu, \sigma^2\} \quad \text{independent}$$

$$P(x_i|w) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{(x_i - \mu)^2}{2\sigma^2}\right]$$

$$\text{Max } P(D|w) \Leftrightarrow \text{Min} -\log P(D|w)$$

$$L = -\log P(D|w) = -\sum_i \log p(x_i|w)$$

$$= \frac{1}{2\sigma^2} \left[\sum_i (x_i - \mu)^2 \right] + N \log \sqrt{2\pi} + \frac{N}{2} \log \sigma^2$$

Reason for logarithm

* Easier derivative

* Avoid underflow.

e.g. $10^{-18} \rightarrow 18$

10

At minimizer.

$$\frac{\partial L}{\partial \mu} = 0 = -\frac{1}{\sigma^2} \sum_i (x_i - \mu)$$

$$\Rightarrow \sum_i (x_i - \mu) = 0 \Rightarrow \mu = \frac{1}{N} \sum_i x_i \rightarrow \text{Statistical mean.}$$

$$\frac{\partial L}{\partial \sigma^2} = 0 = -\frac{1}{2(\sigma^2)^2} \sum_i \frac{(x_i - \mu)^2}{N} + \frac{N}{2} \frac{1}{\sigma^2}$$

$$\Rightarrow \sigma_{ML}^2 = \frac{1}{N} \sum_i (x_i - \mu)^2 \rightarrow \text{Statistical variance (Biased),}$$

$$E[\sigma_{ML}^2] = \frac{N-1}{N} \sigma^2 \neq \sigma^2$$

Reason for bias : μ is dependent on $\{x_i\}$. so degrees-of-freedom of σ_{ML}^2 is $N-1$ instead of N .

Unbiased estimate is $\sigma^2 = \frac{1}{N-1} \sum_i (x_i - \mu)^2 \leftarrow \text{Bessel correction}$

$$\text{Proof: } \sigma_{ML}^2 = \frac{1}{N} \sum_i [(x_i - \mu) - (\mu_{ML} - \mu)]^2 = \frac{1}{N} \left[\sum_i (x_i - \mu)^2 - N(\mu_{ML} - \mu)^2 \right]$$

$$\begin{aligned} E[\sigma_{ML}^2] &= \frac{1}{N} \left[\sum_i E(x_i - \mu)^2 - N E(\mu_{ML} - \mu)^2 \right] \\ &= \frac{1}{N} \left[N \times \sigma^2 - N \cdot \underbrace{\frac{1}{N} \sigma^2}_{\text{II}} \right] \leftarrow \text{Left as exercise} \\ &= \frac{N-1}{N} \sigma^2 \end{aligned}$$

A more Bayesian ML example.

Want to determine a parameter μ in a model

1. Before measurement we have a belief, i.e. prior. $\mu \sim N(\mu_0, \sigma_0^2)$

2. Now measurement of the model is $x \sim N(x_0, \sigma^2)$

The probability of getting x_0 given μ is, the likelihood,

$$p(x_0|\mu) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(\mu-x_0)^2}{2\sigma^2}\right]$$

i.e. when truth is μ , the prob. of getting x_0 in measurement.

3. The measurement updates our belief about μ , i.e. posterior

$$p(\mu|x_0) = \frac{p(x_0|\mu) p(\mu)}{p(x_0)} \propto p(x_0|\mu) p(\mu)$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(\mu-x_0)^2}{2\sigma^2}\right] \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp\left[-\frac{(\mu-\mu_0)^2}{2\sigma_0^2}\right]$$

Cheating; for Gaussian prior & likelihood, posterior is Gaussian.

Assume $p(\mu|x_0) \propto \exp\left[-\frac{(\mu-\mu_1)^2}{2\sigma_1^2}\right]$, only need to determine μ_1, σ_1^2 .

$$\propto \exp\left(-\frac{\mu^2}{2\sigma_1^2} + \frac{\mu\mu_1}{\sigma_1^2}\right)$$

$\exp\left(-\frac{\mu^2}{2\sigma_1^2}\right)$ is const

$$\propto \exp\left[-\frac{\mu^2}{2} \left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2}\right) + \mu \left(\frac{\mu_0}{\sigma_0^2} + \frac{x_0}{\sigma_1^2}\right)\right]$$

$$\text{So } \sigma_1^2 = (\sigma_0^{-2} + \sigma_1^{-2})^{-1}, \quad \mu_1 = \sigma_1^2 (\mu_0 \sigma_0^{-2} + x_0 \sigma_1^{-2})$$

Old guess (μ_0, σ_0^2) updated to (μ_1, σ_1^2) New confidence level.
New guess

Comment: $P(x) = \int p(x|\mu) p(\mu) d\mu$ is called evidence.

$P(x)$ depends on the form of $p(x|\mu)$, i.e. the model.

Good model \leftrightarrow High $p(x)$, hence the name evidence.

or Good measurement in
this particular example